

Limiting Spectral Distribution of Large Sample Covariance Matrices

Marwa BANNA

PhD student at Université Paris-Est Marne-La-Vallée under the direction of
Florence Merlevède and Emmanuel Rio

Young Women in Probability
Bonn, Germany

26 May 2014

Introduction and Motivation

- Let $\mathbf{X}_1, \dots, \mathbf{X}_n \in \mathbb{R}^N$ be i.i.d centered random vectors with covariance matrix $\Sigma = \mathbb{E}(\mathbf{X}_1 \mathbf{X}_1^T) = \dots = \mathbb{E}(\mathbf{X}_n \mathbf{X}_n^T)$.

Introduction and Motivation

- Let $\mathbf{X}_1, \dots, \mathbf{X}_n \in \mathbb{R}^N$ be i.i.d centered random vectors with covariance matrix $\Sigma = \mathbb{E}(\mathbf{X}_1 \mathbf{X}_1^T) = \dots = \mathbb{E}(\mathbf{X}_n \mathbf{X}_n^T)$.
- The sample covariance matrix \mathbf{B}_n is defined by

$$\mathbf{B}_n = \frac{1}{n} \sum_{k=1}^n \mathbf{X}_k \mathbf{X}_k^T = \frac{1}{n} \mathcal{X}_n \mathcal{X}_n^T, \quad \mathcal{X}_n = (\mathbf{X}_1 \dots \mathbf{X}_n) \in \mathcal{M}_{N \times n}(\mathbb{R})$$

Introduction and Motivation

- Let $\mathbf{X}_1, \dots, \mathbf{X}_n \in \mathbb{R}^N$ be i.i.d centered random vectors with covariance matrix $\Sigma = \mathbb{E}(\mathbf{X}_1 \mathbf{X}_1^T) = \dots = \mathbb{E}(\mathbf{X}_n \mathbf{X}_n^T)$.
- The sample covariance matrix \mathbf{B}_n is defined by

$$\mathbf{B}_n = \frac{1}{n} \sum_{k=1}^n \mathbf{X}_k \mathbf{X}_k^T = \frac{1}{n} \mathcal{X}_n \mathcal{X}_n^T, \quad \mathcal{X}_n = (\mathbf{X}_1 \dots \mathbf{X}_n) \in \mathcal{M}_{N \times n}(\mathbb{R})$$

- $\mathbb{E}(\mathbf{B}_n) = \Sigma$
- For fixed N , the strong law of large numbers implies

$$\lim_{n \rightarrow +\infty} \mathbf{B}_n = \lim_{n \rightarrow +\infty} \frac{1}{n} \sum_{k=1}^n \mathbf{X}_k \mathbf{X}_k^T = \Sigma \quad \text{a.s.}$$

Introduction and Motivation

- Let $\mathbf{X}_1, \dots, \mathbf{X}_n \in \mathbb{R}^N$ be i.i.d centered random vectors with covariance matrix $\Sigma = \mathbb{E}(\mathbf{X}_1 \mathbf{X}_1^T) = \dots = \mathbb{E}(\mathbf{X}_n \mathbf{X}_n^T)$.
- The sample covariance matrix \mathbf{B}_n is defined by

$$\mathbf{B}_n = \frac{1}{n} \sum_{k=1}^n \mathbf{X}_k \mathbf{X}_k^T = \frac{1}{n} \mathcal{X}_n \mathcal{X}_n^T, \quad \mathcal{X}_n = (\mathbf{X}_1 \dots \mathbf{X}_n) \in \mathcal{M}_{N \times n}(\mathbb{R})$$

- $\mathbb{E}(\mathbf{B}_n) = \Sigma$
- For fixed N , the strong law of large numbers implies

$$\lim_{n \rightarrow +\infty} \mathbf{B}_n = \lim_{n \rightarrow +\infty} \frac{1}{n} \sum_{k=1}^n \mathbf{X}_k \mathbf{X}_k^T = \Sigma \quad \text{a.s.}$$

- What happens once $N := N_n \rightarrow \infty$ as $n \rightarrow \infty$?

Marčenko-Pastur Theorem

- Let $(X_{ij})_{i,j \geq 1}$ be a family of i.i.d. random variables such that $\mathbb{E}(X_{11}) = 0$ and $\text{Var}(X_{11}) = \sigma^2$.
- Let $\mathbf{B}_n = \frac{1}{n} \sum_{k=1}^n \mathbf{X}_k \mathbf{X}_k^T = \frac{1}{n} \mathcal{X}_n \mathcal{X}_n^T$ where

$$\mathcal{X}_n = (\mathbf{X}_1 \dots \mathbf{X}_n) = \begin{pmatrix} X_{11} & X_{12} & \dots & X_{1n} \\ \vdots & \vdots & & \vdots \\ X_{N1} & X_{N2} & \dots & X_{Nn} \end{pmatrix}.$$

Marčenko-Pastur Theorem

- Let $(X_{ij})_{i,j \geq 1}$ be a family of i.i.d. random variables such that $\mathbb{E}(X_{11}) = 0$ and $\text{Var}(X_{11}) = \sigma^2$.
- Let $\mathbf{B}_n = \frac{1}{n} \sum_{k=1}^n \mathbf{X}_k \mathbf{X}_k^T = \frac{1}{n} \mathcal{X}_n \mathcal{X}_n^T$ where

$$\mathcal{X}_n = (\mathbf{X}_1 \dots \mathbf{X}_n) = \begin{pmatrix} X_{11} & X_{12} & \dots & X_{1n} \\ \vdots & \vdots & & \vdots \\ X_{N1} & X_{N2} & \dots & X_{Nn} \end{pmatrix}.$$

- The empirical spectral measure of \mathbf{B}_n is defined by

$$\mu_{\mathbf{B}_n} = \frac{1}{N} \sum_{k=1}^N \delta_{\lambda_k}$$

where $\lambda_1, \dots, \lambda_N$ are the eigenvalues of \mathbf{B}_n .

Marčenko-Pastur Theorem

- Let $(X_{ij})_{i,j \geq 1}$ be a family of i.i.d. random variables such that $\mathbb{E}(X_{11}) = 0$ and $\text{Var}(X_{11}) = \sigma^2$.
- Let $\mathbf{B}_n = \frac{1}{n} \sum_{k=1}^n \mathbf{X}_k \mathbf{X}_k^T = \frac{1}{n} \mathcal{X}_n \mathcal{X}_n^T$ where

$$\mathcal{X}_n = (\mathbf{X}_1 \dots \mathbf{X}_n) = \begin{pmatrix} X_{11} & X_{12} & \dots & X_{1n} \\ \vdots & \vdots & & \vdots \\ X_{N1} & X_{N2} & \dots & X_{Nn} \end{pmatrix}.$$

- The empirical spectral measure of \mathbf{B}_n is defined by

$$\mu_{\mathbf{B}_n} = \frac{1}{N} \sum_{k=1}^N \delta_{\lambda_k}$$

where $\lambda_1, \dots, \lambda_N$ are the eigenvalues of \mathbf{B}_n .

- We suppose that $c_n := \frac{N}{n} \xrightarrow{n \rightarrow +\infty} c \in (0, \infty)$.

Marčenko-Pastur Theorem

Theorem

For any continuous and bounded function $f : \mathbb{R} \rightarrow \mathbb{R}$,

$$\int f d\mu_{\mathbf{B}_n} \xrightarrow{n \rightarrow +\infty} \int f d\mu_c \quad \text{a.s.}$$

where μ_c is the Marčenko-Pastur law

$$\left(1 - \frac{1}{c}\right)_+ \delta_0 + \frac{1}{2\pi c \sigma^2 x} \sqrt{(b-x)(x-a)} \mathbf{1}_{[a,b]}(x) dx$$

with $\cdot_+ := \max(0, \cdot)$, $a = \sigma^2(1 - \sqrt{c})^2$ and $b = \sigma^2(1 + \sqrt{c})^2$.

The Stieltjes Transform

The Stieltjes transform $S_G : \mathbb{C}_+ \rightarrow \mathbb{C}$ of a measure ν on \mathbb{R} is defined by

$$S_\nu(z) := \int \frac{1}{x - z} d\nu(x)$$

The Stieltjes Transform

The Stieltjes transform $S_G : \mathbb{C}_+ \rightarrow \mathbb{C}$ of a measure ν on \mathbb{R} is defined by

$$S_\nu(z) := \int \frac{1}{x - z} d\nu(x)$$

- $|S_\nu(z)| \leq 1/\Im(z)$ and $\Im(S_\nu(z)) \geq 0$

The Stieltjes Transform

The Stieltjes transform $S_G : \mathbb{C}_+ \rightarrow \mathbb{C}$ of a measure ν on \mathbb{R} is defined by

$$S_\nu(z) := \int \frac{1}{x - z} d\nu(x)$$

- $|S_\nu(z)| \leq 1/\Im(z)$ and $\Im(S_\nu(z)) \geq 0$
- The function S_ν is analytic over \mathbb{C}_+ and characterizes ν

$$\nu([a, b]) = \lim_{y \downarrow 0} \frac{1}{\pi} \int_a^b \Im S_\nu(x + iy) dx$$

The Stieltjes Transform

The Stieltjes transform $S_G : \mathbb{C}_+ \rightarrow \mathbb{C}$ of a measure ν on \mathbb{R} is defined by

$$S_\nu(z) := \int \frac{1}{x - z} d\nu(x)$$

- $|S_\nu(z)| \leq 1/\Im(z)$ and $\Im(S_\nu(z)) \geq 0$
- The function S_ν is analytic over \mathbb{C}_+ and characterizes ν

$$\nu([a, b]) = \lim_{y \downarrow 0} \frac{1}{\pi} \int_a^b \Im S_\nu(x + iy) dx$$

- For a sequence of measures $(\nu_n)_n$ on \mathbb{R} , we have

$$\left(\nu_n \xrightarrow[n \rightarrow \infty]{\mathcal{L}} \nu \right) \Leftrightarrow \left(\forall z \in \mathbb{C}_+, S_{\nu_n}(z) \xrightarrow[n \rightarrow \infty]{} S_\nu(z) \right).$$

Non linear Stationary Process

- $X = (X_k)_{k \in \mathbb{Z}}$ is said to be stationary if $\forall k \in \mathbb{Z}$ and $\forall \ell \in \mathbb{N}$,
$$(X_k, \dots, X_{k+\ell}) \stackrel{\mathcal{D}}{\sim} (X_0, \dots, X_\ell)$$

Non linear Stationary Process

- Let $(\varepsilon_k)_{k \in \mathbb{Z}}$ be a sequence of i.i.d. random variables

Non linear Stationary Process

- Let $(\varepsilon_k)_{k \in \mathbb{Z}}$ be a sequence of i.i.d. random variables
- Let $g : \mathbb{R}^{\mathbb{Z}} \rightarrow \mathbb{R}$ be a measurable function such that for any $k \in \mathbb{Z}$,

$$X_k = g(\xi_k) \text{ with } \xi_k = (\dots, \varepsilon_{k-1}, \varepsilon_k)$$

is well-defined, $\mathbb{E}(X_k) = 0$ and $\|X_k\|_2 < \infty$.

Non linear Stationary Process

- Let $(\varepsilon_k)_{k \in \mathbb{Z}}$ be a sequence of i.i.d. random variables
- Let $g : \mathbb{R}^{\mathbb{Z}} \rightarrow \mathbb{R}$ be a measurable function such that for any $k \in \mathbb{Z}$,

$$X_k = g(\xi_k) \text{ with } \xi_k = (\dots, \varepsilon_{k-1}, \varepsilon_k)$$

is well-defined, $\mathbb{E}(X_k) = 0$ and $\|X_k\|_2 < \infty$.

- For $i = 1, \dots, n$, let $(X_k^i)_{k \in \mathbb{Z}}$ be n independent copies of $(X_k)_{k \in \mathbb{Z}}$

Non linear Stationary Process

- Let $(\varepsilon_k)_{k \in \mathbb{Z}}$ be a sequence of i.i.d. random variables
- Let $g : \mathbb{R}^{\mathbb{Z}} \rightarrow \mathbb{R}$ be a measurable function such that for any $k \in \mathbb{Z}$,

$$X_k = g(\xi_k) \text{ with } \xi_k = (\dots, \varepsilon_{k-1}, \varepsilon_k)$$

is well-defined, $\mathbb{E}(X_k) = 0$ and $\|X_k\|_2 < \infty$.

- For $i = 1, \dots, n$, let $(X_k^i)_{k \in \mathbb{Z}}$ be n independent copies of $(X_k)_{k \in \mathbb{Z}}$

$$\mathcal{X}_n = \begin{pmatrix} X_1^1 & X_1^2 & \dots & X_1^n \\ \vdots & \vdots & & \vdots \\ X_N^1 & X_N^2 & \dots & X_N^n \end{pmatrix}$$

$\mathbf{X}_1 \quad \mathbf{X}_2 \quad \mathbf{X}_n$

Non linear Stationary Process

- Let $(\varepsilon_k)_{k \in \mathbb{Z}}$ be a sequence of i.i.d. random variables
- Let $g : \mathbb{R}^{\mathbb{Z}} \rightarrow \mathbb{R}$ be a measurable function such that for any $k \in \mathbb{Z}$,

$$X_k = g(\xi_k) \text{ with } \xi_k = (\dots, \varepsilon_{k-1}, \varepsilon_k)$$

is well-defined, $\mathbb{E}(X_k) = 0$ and $\|X_k\|_2 < \infty$.

- For $i = 1, \dots, n$, let $(X_k^i)_{k \in \mathbb{Z}}$ be n independent copies of $(X_k)_{k \in \mathbb{Z}}$

$$\mathcal{X}_n = \begin{pmatrix} X_1^1 & X_1^2 & \dots & X_1^n \\ \vdots & \vdots & & \vdots \\ X_N^1 & X_N^2 & \dots & X_N^n \end{pmatrix}$$

$\mathbf{X}_1 \quad \mathbf{X}_2 \quad \dots \quad \mathbf{X}_n$

- $\mathbf{B}_n = \frac{1}{n} \mathcal{X}_n \mathcal{X}_n^T = \frac{1}{n} \sum_{k=1}^n \mathbf{X}_k \mathbf{X}_k^T$

Theorem: Banna and Merlevède (2013)

Suppose that $\lim_{n \rightarrow \infty} N/n = c \in (0, \infty)$ and

$$\sum_{k \geq 0} \|P_0(X_k)\|_2 < \infty \text{ where } P_0(X_k) = \mathbb{E}(X_k|\xi_0) - \mathbb{E}(X_k|\xi_{-1})$$

Then with probability one $\mu_{\mathbf{B}_n}$ converges weakly to a non-random probability measure whose Stieltjes transform $S = S(z)$ satisfies the equation

$$z = -\frac{1}{\underline{S}} + \frac{c}{2\pi} \int_0^{2\pi} \frac{1}{\underline{S} + (2\pi f(\lambda))^{-1}} d\lambda,$$

where $\underline{S}(z) := -(1-c)/z + cS(z)$ and $f(\cdot)$ is the spectral density of $(X_k)_{k \in \mathbb{Z}}$.

Theorem: Banna and Merlevède (2013)

Suppose that $\lim_{n \rightarrow \infty} N/n = c \in (0, \infty)$ and

$$\sum_{k \geq 0} \|P_0(X_k)\|_2 < \infty \text{ where } P_0(X_k) = \mathbb{E}(X_k|\xi_0) - \mathbb{E}(X_k|\xi_{-1})$$

Then with probability one $\mu_{\mathbf{B}_n}$ converges weakly to a non-random probability measure whose Stieltjes transform $S = S(z)$ satisfies the equation

$$z = -\frac{1}{\underline{S}} + \frac{c}{2\pi} \int_0^{2\pi} \frac{1}{\underline{S} + (2\pi f(\lambda))^{-1}} d\lambda,$$

where $\underline{S}(z) := -(1-c)/z + cS(z)$ and $f(\cdot)$ is the spectral density of $(X_k)_{k \in \mathbb{Z}}$.

$$f(x) = \frac{1}{2\pi} \sum_k \text{Cov}(X_0, X_k) e^{ixk}, \quad x \in \mathbb{R}$$

Applications

- $X_k = \sum_{i \geq 0} a_i \varepsilon_{k-i}$ where $(\varepsilon_i)_{i \in \mathbb{Z}}$ is a sequence of i.i.d. centered random variables then

$$\mathbb{E}(X_k | \xi_0) = \sum_{i \geq k} a_i \varepsilon_{k-i} \quad \text{and} \quad \mathbb{E}(X_k | \xi_{-1}) = \sum_{i \geq k+1} a_i \varepsilon_{k-i}$$

Thus $P_0(X_k) = a_k \varepsilon_0$.

Applications

- $X_k = \sum_{i \geq 0} a_i \varepsilon_{k-i}$ where $(\varepsilon_i)_{i \in \mathbb{Z}}$ is a sequence of i.i.d. centered random variables then

$$\mathbb{E}(X_k | \xi_0) = \sum_{i \geq k} a_i \varepsilon_{k-i} \quad \text{and} \quad \mathbb{E}(X_k | \xi_{-1}) = \sum_{i \geq k+1} a_i \varepsilon_{k-i}$$

Thus $P_0(X_k) = a_k \varepsilon_0$. The condition $\sum_{k \geq 0} \|P_0(X_k)\|_2 < \infty$ is satisfied once

$$\sum_{k \geq 0} |a_k| < \infty \quad \text{and} \quad \|\varepsilon_0\|_2 < \infty.$$

Applications

- $X_k = \sum_{i \geq 0} a_i \varepsilon_{k-i}$ where $(\varepsilon_i)_{i \in \mathbb{Z}}$ is a sequence of i.i.d. centered random variables then

$$\mathbb{E}(X_k | \xi_0) = \sum_{i \geq k} a_i \varepsilon_{k-i} \quad \text{and} \quad \mathbb{E}(X_k | \xi_{-1}) = \sum_{i \geq k+1} a_i \varepsilon_{k-i}$$

Thus $P_0(X_k) = a_k \varepsilon_0$. The condition $\sum_{k \geq 0} \|P_0(X_k)\|_2 < \infty$ is satisfied once

$$\sum_{k \geq 0} |a_k| < \infty \quad \text{and} \quad \|\varepsilon_0\|_2 < \infty.$$

- Functions of linear processes (ex: Riesz-Raikov sums)
- ARCH models

Strategy of Proof

Our aim: $\lim_{n \rightarrow \infty} S_{\mu_{\mathbf{B}_n}}(z) = S(z), \quad \forall z \in \mathbb{C}^+.$

1. $S_{\mu_{\mathbf{B}_n}}(z) - \mathbb{E}(S_{\mu_{\mathbf{B}_n}}(z)) \rightarrow 0$ a.s.

Strategy of Proof

Our aim: $\lim_{n \rightarrow \infty} S_{\mu_{\mathbf{B}_n}}(z) = S(z), \quad \forall z \in \mathbb{C}^+.$

1. $S_{\mu_{\mathbf{B}_n}}(z) - \mathbb{E}(S_{\mu_{\mathbf{B}_n}}(z)) \rightarrow 0$ a.s.
2. $\mathbb{E}(S_{\mu_{\mathbf{B}_n}}(z)) - \mathbb{E}(S_{\mu_{\mathbf{G}_n}}(z)) \rightarrow 0$ where $\mathbf{G}_n = \frac{1}{n} \sum_{k=1}^n \mathbf{Z}_k \mathbf{Z}_k^T$

$$\mathcal{Z}_n = \begin{pmatrix} Z_1^1 & Z_1^2 & \cdots & Z_1^n \\ \vdots & \vdots & & \vdots \\ Z_N^1 & Z_N^2 & \cdots & Z_N^n \\ \mathbf{Z}_1 & \mathbf{Z}_2 & & \mathbf{Z}_n \end{pmatrix}$$

$((Z_k^i)_{k \in \mathbb{Z}})_{1 \leq i \leq n}$ are n independent copies of $(Z_k)_{k \in \mathbb{Z}}$ such that for any $i, j \in \mathbb{Z}$,

$$\text{Cov}(Z_i, Z_j) = \text{Cov}(X_i, X_j)$$

Strategy of Proof

Our aim: $\lim_{n \rightarrow \infty} S_{\mu_{\mathbf{B}_n}}(z) = S(z), \quad \forall z \in \mathbb{C}^+.$

1. $S_{\mu_{\mathbf{B}_n}}(z) - \mathbb{E}(S_{\mu_{\mathbf{B}_n}}(z)) \rightarrow 0$ a.s.
2. $\mathbb{E}(S_{\mu_{\mathbf{B}_n}}(z)) - \mathbb{E}(S_{\mu_{\mathbf{G}_n}}(z)) \rightarrow 0$ where $\mathbf{G}_n = \frac{1}{n} \sum_{k=1}^n \mathbf{Z}_k \mathbf{Z}_k^T$

$$\mathcal{Z}_n = \begin{pmatrix} Z_1^1 & Z_1^2 & \cdots & Z_1^n \\ \vdots & \vdots & & \vdots \\ Z_N^1 & Z_N^2 & \cdots & Z_N^n \\ \mathbf{Z}_1 & \mathbf{Z}_2 & & \mathbf{Z}_n \end{pmatrix}$$

$((Z_k^i)_{k \in \mathbb{Z}})_{1 \leq i \leq n}$ are n independent copies of $(Z_k)_{k \in \mathbb{Z}}$ such that for any $i, j \in \mathbb{Z}$,

$$\text{Cov}(Z_i, Z_j) = \text{Cov}(X_i, X_j)$$

3. $\mathbb{E}(S_{\mu_{\mathbf{G}_n}}(z)) - S(z) \rightarrow 0$

Strategy of Proof

Our aim: $\lim_{n \rightarrow \infty} S_{\mu_{\mathbf{B}_n}}(z) = S(z), \quad \forall z \in \mathbb{C}^+.$

1. $S_{\mu_{\mathbf{B}_n}}(z) - \mathbb{E}(S_{\mu_{\mathbf{B}_n}}(z)) \rightarrow 0$ a.s. **Concentration inequality**
2. $\mathbb{E}(S_{\mu_{\mathbf{B}_n}}(z)) - \mathbb{E}(S_{\mu_{\mathbf{G}_n}}(z)) \rightarrow 0$ where $\mathbf{G}_n = \frac{1}{n} \sum_{k=1}^n \mathbf{Z}_k \mathbf{Z}_k^T$

$$\mathcal{Z}_n = \begin{pmatrix} Z_1^1 & Z_1^2 & \cdots & Z_1^n \\ \vdots & \vdots & & \vdots \\ Z_N^1 & Z_N^2 & \cdots & Z_N^n \\ \mathbf{Z}_1 & \mathbf{Z}_2 & & \mathbf{Z}_n \end{pmatrix}$$

$((Z_k^i)_{k \in \mathbb{Z}})_{1 \leq i \leq n}$ are n independent copies of $(Z_k)_{k \in \mathbb{Z}}$ such that for any $i, j \in \mathbb{Z}$,

$$\text{Cov}(Z_i, Z_j) = \text{Cov}(X_i, X_j)$$

3. $\mathbb{E}(S_{\mu_{\mathbf{G}_n}}(z)) - S(z) \rightarrow 0$

Strategy of Proof

Our aim: $\lim_{n \rightarrow \infty} S_{\mu_{\mathbf{B}_n}}(z) = S(z), \quad \forall z \in \mathbb{C}^+.$

1. $S_{\mu_{\mathbf{B}_n}}(z) - \mathbb{E}(S_{\mu_{\mathbf{B}_n}}(z)) \rightarrow 0$ a.s. **Concentration inequality**
2. $\mathbb{E}(S_{\mu_{\mathbf{B}_n}}(z)) - \mathbb{E}(S_{\mu_{\mathbf{G}_n}}(z)) \rightarrow 0$ where $\mathbf{G}_n = \frac{1}{n} \sum_{k=1}^n \mathbf{Z}_k \mathbf{Z}_k^T$

$$\mathcal{Z}_n = \begin{pmatrix} Z_1^1 & Z_1^2 & \cdots & Z_1^n \\ \vdots & \vdots & & \vdots \\ Z_N^1 & Z_N^2 & \cdots & Z_N^n \\ \mathbf{Z}_1 & \mathbf{Z}_2 & & \mathbf{Z}_n \end{pmatrix}$$

$((Z_k^i)_{k \in \mathbb{Z}})_{1 \leq i \leq n}$ are n independent copies of $(Z_k)_{k \in \mathbb{Z}}$ such that for any $i, j \in \mathbb{Z}$,

$$\text{Cov}(Z_i, Z_j) = \text{Cov}(X_i, X_j)$$

3. $\mathbb{E}(S_{\mu_{\mathbf{G}_n}}(z)) - S(z) \rightarrow 0$ **Silverstein (1995)**

Strategy of Proof

Our aim: $\lim_{n \rightarrow \infty} S_{\mu_{\mathbf{B}_n}}(z) = S(z), \quad \forall z \in \mathbb{C}^+.$

1. $S_{\mu_{\mathbf{B}_n}}(z) - \mathbb{E}(S_{\mu_{\mathbf{B}_n}}(z)) \rightarrow 0$ a.s. **Concentration inequality**
2. $\mathbb{E}(S_{\mu_{\mathbf{B}_n}}(z)) - \mathbb{E}(S_{\mu_{\mathbf{G}_n}}(z)) \rightarrow 0$ where $\mathbf{G}_n = \frac{1}{n} \sum_{k=1}^n \mathbf{Z}_k \mathbf{Z}_k^T$




$$\mathcal{Z}_n = \begin{pmatrix} Z_1^1 & Z_1^2 & \cdots & Z_1^n \\ \vdots & \vdots & & \vdots \\ Z_N^1 & Z_N^2 & \cdots & Z_N^n \\ \mathbf{Z}_1 & \mathbf{Z}_2 & & \mathbf{Z}_n \end{pmatrix}$$

$((Z_k^i)_{k \in \mathbb{Z}})_{1 \leq i \leq n}$ are n independent copies of $(Z_k)_{k \in \mathbb{Z}}$ such that for any $i, j \in \mathbb{Z}$,

$$\text{Cov}(Z_i, Z_j) = \text{Cov}(X_i, X_j)$$

Technic of blocks associated with the Lindeberg method

3. $\mathbb{E}(S_{\mu_{\mathbf{G}_n}}(z)) - S(z) \rightarrow 0$ **Silverstein (1995)**

-  Banna, M., Merlevede, F. (2013). Limiting spectral distribution of large sample covariance matrices associated with a class of stationary processes. *Journal of Theoretical Probability*, 1-39.
-  Chatterjee, S. (2006). A generalization of the Lindeberg principle. *Ann. Probab.* **34**, 2061-2076.
-  Marčenko, V. and Pastur, L. (1967). Distribution of eigenvalues for some sets of random matrices. *Mat. Sb.* **72**, 507-536.

Thank you for your attention!